

# Economic repercussions of fisheries-induced evolution

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**Fish stocks experiencing high fishing mortality show a tendency to mature earlier and at a smaller size, which may have a genetic component and therefore long-lasting economic and biological effects. To date, the economic effects of such ecoevolutionary dynamics have not been empirically investigated. Using 70 y of data, we develop a bioeconomic model for Northeast Arctic cod to compare the economic yield in a model in which life-history traits can vary only through phenotypic plasticity with a model in which, in addition, genetic changes can occur. We find that evolutionary changes toward faster growth and earlier maturation occur consistently even if a stock is optimally managed. However, if a stock is managed optimally, the evolutionary changes actually increase economic yield because faster growth and earlier maturation raise the stock's productivity. The optimal fishing mortality is almost identical for the evolutionary and nonevolutionary model and substantially lower than what it has been historically. Therefore, the costs of ignoring evolution under optimal management regimes are negligible. However, if fishing mortality is as high as it has been historically, evolutionary changes may result in economic losses, but only if the fishery is selecting for medium-sized individuals. Because evolution facilitates growth, the fish are younger and still immature when they are susceptible to getting caught, which outweighs the increase in productivity due to fish spawning at an earlier age.**

Atlantic cod | genetic adaptations | harvest control rule | marine governance | adaptive management

Life-history theory, experiments, and field-based studies strongly suggest that fishing is capable of inducing genetic adaptations, especially when it removes individuals with characteristics such as large body size (1–5). Even if fishing is not size-selective, high fishing mortality may be sufficient to induce genetic change (6, 7). It is difficult to predict how genetic changes at the individual level affect population-level properties. Genetic adaptations may, in principle, be beneficial for the state of a stock, by enabling individuals to invest more into reproduction and growth (1, 8). As a consequence, the stock may become more productive, allowing exploited populations to withstand higher fishing mortalities than they could in the absence of such adaptation, possibly permitting higher yields. However, although an individual's increased reproductive investment leads to larger gonads, this happens at the expense of slower postmaturation growth. Maturing earlier may also reduce fecundity, because individuals are smaller when they reproduce (9). Moreover, adapting to fishing may bear a cost of maladaptation, resulting in increased natural mortality (10, 11). Therefore, fisheries-induced evolution (FIE) may reduce yield (2, 4, 12, 13) and may even imply a “Darwinian debt” (14) to be paid back by future generations, at least if genetic changes are difficult to reverse (1, 15, 16). Clearly, FIE has the potential for causing positive and negative effects on key stock properties such as spawning stock biomass (SSB) and yield, making its economic effect ambiguous. It is also an open question whether the expected size of the economic effects are substantial, largely because any evolutionary changes are closely intertwined with ecological effects. For example, the release of density dependence when population biomass is fished down could be an important driver of phenotypic

change (1, 17, 18) and might override effects of FIE on yield. However, the economic consequences of FIE and its effects on optimal fishing mortalities in wild populations have yet to be determined. Here, we ask how evolutionarily informed management differs from classical fisheries management. First, we determine how an evolving fish population should be optimally managed. Second, we analyze how these management strategies differ compared with optimal management derived for a population whose development is purely determined by ecological processes. Third, we ask how substantial the losses are if a fishery's manager—unaware of any evolutionary changes—manages an evolving population as if it were not evolving. Fourth, we analyze how FIE affects the performance of the fishery that is not optimally managed, but heavily exploited.

Northeast Arctic (NEA) cod is currently the world's largest stock of Atlantic cod (*Gadus morhua*) and provides substantial ecosystem services. The stock's fishery is an important economic resource for Norway and Russia, with annual catches by Norway being worth more than 500 million US dollars in 2010, and Russia obtaining about the same revenue. Traditionally, harvesting focused on adult cod at the stock's spawning grounds along the Norwegian coast. From the 1930s, when industrial trawlers were introduced in the stock's feeding grounds in the Barents Sea, immature fish came under substantial fishing pressure, and total fishing mortality increased (19). Evolutionary changes have been predicted to be a factor in explaining the observed declines in age and length at maturation in NEA cod, although the predicted extent has varied among studies (17, 20, 21).

We develop a bioeconomic model to investigate if and how FIE affects economic yield (Fig. 1). Our model is a comprehensive compilation of a life-history model for a harvested species with economic components relying on individual vessel data, making this a unique empirically derived bioeconomic model for investigating genetic adaptations to harvesting. This model has been specifically built for NEA cod to investigate the ecological and evolutionary effects of exploitation on the changes in maturation that occurred after fishing mortality was intensified in the 1930s in the feeding grounds (17). To match the observed trends in the biological model as closely as possible, we recreated the historical selection pressure to determine the evolvability (i.e., the coefficient of genetic variation) in the life-history traits (17). Although we focus on the feeding ground fishery in the Barents Sea, we also included fishing in the spawning grounds at the historic levels between 1932 until 2005, and at a constant rate after 2006. Hence, we consider the spawning ground fishery to

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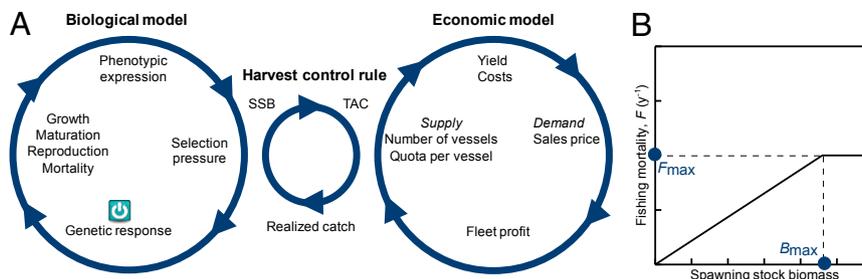
The authors declare no conflict of interest.

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**Fig. 1.** An overview of the bioeconomic model. (A) The biological and economic model components are coupled by the harvest control rule (HCR). The individual-based biological model describes the evolution of key life-history traits if genetic changes are allowed to occur in the model. The economic model accounts for the supply and demand side of the fishery, as well as for fleet profit generated. (B) The shape of the HCR depends on two parameters: above the level  $B_{\max}$  of spawning stock biomass, the maximum fishing mortality  $F_{\max}$  is allowed. Between  $B_{\max}$  and a biomass level of zero, fishing mortality linearly decreases from  $F_{\max}$  to zero. The structure of this HCR is in agreement with that advised in 2004 by the International Council for the Exploration of the Sea for the Northeast Arctic cod fishery.

be beyond the control of the manager. The biological model component is built on the individual-based ecogenetic model framework developed by Dunlop et al. (1), describing four evolving life-history traits capturing key aspects of growth, maturation, and reproduction (Table S1). Changes in life-history traits may be driven both by ecological processes, such as phenotypic plasticity and density dependence, and genetic processes. To evaluate whether accounting for FIE requires a special harvest strategy, we also analyze a nonevolutionary version of the biological model in which the genetic traits cannot evolve. We therefore compare a nonevolutionary model, in which changes in populations are driven only by phenotypic plasticity, with an evolutionary model that allows, in addition, for genetic adaptations. The economic model component consists of production and cost functions estimated specifically for the Norwegian cod trawler fleet (Table S1). We incorporate a demand function, also estimated from empirical data, to account for how total catch affects the price of landings (22). Our model incorporates feedbacks between the stock development and the economic gains through an optimal harvest control rule (HCR), which is constrained by the two parameters  $B_{\max}$  and  $F_{\max}$  (Fig. 1). This shape makes it directly comparable to the HCR that is being implemented for NEA cod since 2004 (23, 24). We search for the parameter combination that gives the highest net present value (NPV) of fleet profits. We derive HCRs that are optimized in either the evolutionary or nonevolutionary versions of the model.

## Results

We first compare the emerging properties of the evolutionary model with the nonevolutionary model, when both are managed according to what an HCR recommends that has been optimized for fleet profits (Table 1, Evolution vs. Ecology). We find that the optimal fishing mortality is almost identical for the evolutionary and nonevolutionary models and substantially lower than what it has been historically. Despite this, the emerging biomass

**Table 1. Optimal HCR for the evolutionary model (evolution) and nonevolutionary model (ecology)**

Model	$F$	TAC	SSB	NPV
Evolution	0.34	469 (60)	767 (163)	25.4
Ecology	0.35	443 (48)	643 (118)	25.3
Evolution ignored	0.35	470 (60)	735 (155)	25.4

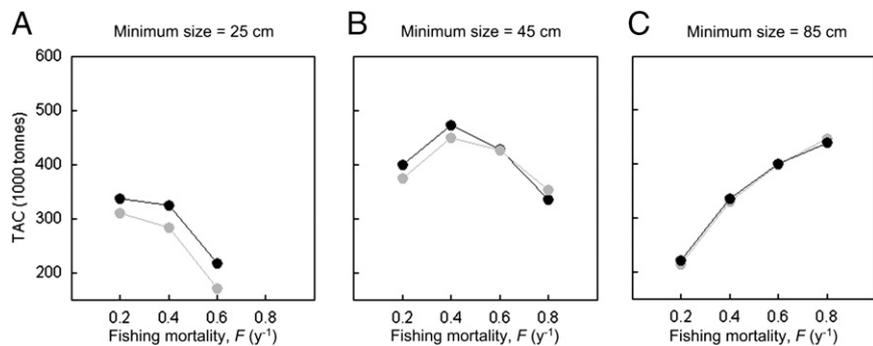
Values shown are averages for 1932–2100 of fishing mortality ( $F$ ), total allowable catch (TAC), spawning stock biomass (SSB), with temporal SDs in parentheses, and net present value (NPV) is given for a discount rate of 2%. “Evolution ignored” uses the evolutionary model with the ecologically optimal harvest control rule (HCR). Units:  $F$  ( $y^{-1}$ ), TAC and SSB (1,000 tonnes), NPV (in billions, US dollars).

levels and the total allowable catch (TAC) are higher in the evolutionary model, indicating that evolution indeed makes the stock more productive, permitting higher yields for the same fishing mortality. Overall, the NPV of the fishery is higher when evolution occurs, even though the total effect is very small. Given that the recommended fishing mortalities are almost identical, the loss of disregarding any evolutionary effects is negligible, and the NPV is still higher if evolution occurs and is ignored by managers (Table 1, Evolution ignored). The key message here is that a low fishing mortality is optimal, regardless of whether genetic changes occur. This prediction holds for different discount rates (Table S2), when sales prices are assumed to be independent of the total catch, and when the price that can be obtained per kilogram of cod rises with the weight of the fish (Table S3).

Given that fishing mortality has not been low for the NEA cod fishery in the past, and considering that worldwide most fisheries are still far from being managed optimally, we also investigate how evolution affects the stock when it is overexploited. To do so, we use historic fishing mortalities between 1932 and 2006, and the average fishing mortality afterwards to simulate a scenario of high fishing pressure; this is then contrasted with a counterfactual scenario that analyzes how the fate of the fishery would have developed if an optimal HCR had been introduced already in 1932 (as given in Table 1, Evolution). We find that using an optimal HCR leads to higher biomass levels in the evolutionary model, compared with the case where only ecological effects are present. The opposite is true for the scenario of historically high fishing mortality, where biomass is slightly lower in the evolutionary model (Fig. 2A). As a result, the corresponding TAC and NPV are also slightly lower when evolution occurs and fishing mortality is high (Table S4).

It is not immediately obvious why evolution has a positive effect on the fishery if fishing mortality is set optimally, but a negative effect if fishing mortality is high. Inspecting key life-history traits reveals that age at maturation declines over time in all scenarios (Fig. 2B), and although this also occurs in the nonevolutionary model (solely as a result of phenotypic plasticity), the decline is even more severe when evolution takes place. A decline in length at maturation occurs in all scenarios as well, and is even more pronounced if fishing mortality is high (Fig. 2C). Despite reduced age and length at maturation, the reproductive output per unit of SSB, a measure of the stock's productivity, is increasing over time when evolution occurs (Fig. 2D). To better understand the population structure, we take a closer look at the age composition at the simulation endpoints (Fig. 3). We find that despite individual fish being smaller at maturation, the size at a given age is consistently larger for the evolutionary model compared with the nonevolutionary model, irrespective of the fishing mortality being optimal or high (Fig. 3A). Indeed, the underlying genetic trait changes show that the evolving population invests more in intrinsic somatic growth





**Fig. 4.** (A–C) Total allowable catch (TAC) under different minimum size limits and for different constant fishing mortalities,  $F$ . The evolutionary model (black) predicts higher TAC than the nonevolutionary model (gray) when selection also acts on very young fish. For a minimum size limit of 85 cm, the two model predictions become essentially indistinguishable. At the intermediate minimum size limit of 45 cm, the TAC is highest for the evolutionary model when fishing mortality is low, but as fishing intensity increases, the TAC is smaller for the evolutionary model.

overridden by making those fish more vulnerable that are larger, but also younger and still immature (Table S5).

In this study, the coefficient of genetic variation was set at a level that resulted in the best fit to empirical observations in age and length at maturation (Table S1), but we nonetheless investigated the effect of this parameter (the evolvability of traits) on model predictions. As expected (1, 16, 25, 26), higher genetic variance resulted in fish maturing at even younger ages and smaller sizes, while also growing faster. Consequently, higher TACs can be obtained when the evolvability is high, suggesting that stronger evolutionary forces can have a positive effect on the fishery (Fig. S3).

## Discussion

Our model predicts that evolutionary change occurs even if fishing mortality is low, which implies that a management strategy aimed at avoiding genetic change might not be feasible. At the same time, we find that fisheries-induced evolution is not necessarily bad for the fishery, and most of the time even beneficial—especially a fishery that is managed according to what is ecologically optimal can safely ignore any evolutionary effects, at least for the stock and under the conditions that we are considering. This finding is very surprising and in contrast to much of the existing literature, which tends to sketch a gloomy picture of the potential consequences of FIE. It is also comforting that fishing can cause evolution of faster growth, allowing the population to withstand higher harvest pressure and prevent stock collapse (Fig. S3). Nonetheless, the life-history changes we predict could have management implications because they affect important indicators that are commonly used to assess the state of the stock. Evolution tends to increase the ratio between SSB and total biomass (Fig. S4), which could mask a decreasing trend in total biomass and affect the stock-recruitment relationship (26); this may have important management implications when biomass levels approach SSB-based limit reference points (27, 28). Even more worrisome is our finding that evolutionary effects tend to be more important when a fish stock is overexploited and the fishery is intermediately size-selective. Admittedly, such an exploitation regime is a special case, but unfortunately the one that, worldwide, most fisheries are facing. Surprisingly, an economic cost of evolution under these conditions does not materialize because of a drop in reproductive output or, as many might expect, because of a reduction in growth or size-at-age (29). On the contrary, evolution here promoted faster growth, yet still could exact an economic cost. These results underscore the importance of management taking into account the detailed age and size structure of the stock (30–32).

Although we find that selectively removing individuals of intermediate size may result in economic losses due to evolutionary

change, we do not find any evidence that targeting only large fish results in evolutionary loss (Fig. 4). These findings may shed light on the discussion whether harvesting should be balanced or selective (33). In this study, we assume a knife-edge selectivity in our model (34, 35), so different gear types with other selectivity patterns remain to be explored by further research. Although gear regulations can, in principle, be easily changed, our findings may also hint at broader problems. If predation is size-selective, evolutionary changes may affect natural mortality which may lead to similar consequences as fishing mortality (10, 11). Investigating how FIE acts in concert with natural mortality, climatic changes, or other driving forces remains to be explored, especially in the light of recovery potential (16).

Although our biological model is complex, the optimal HCR was constrained by two parameters, resembling the shape of the HCR currently adopted for NEA cod. It would be interesting to see to what extent our results carry over to a simpler biological model that could then be used for more flexible optimization routines treating the minimum size limit, for example, as a choice variable. Another interesting avenue is to separately optimize HCRs for the NEA cod's feeding and spawning grounds. Previous research has found predictions for FIE to differ depending on whether management actions target feeding or spawning grounds (36). Here, we focused on the fishery in the stock's feeding grounds and kept the fishing mortality at observed levels in the stock's spawning grounds to mimic the historic selection pressure on mature fish, while parsimoniously asking what can be changed for the trawler fleet in the Barents Sea.

Together, our results show that the economic consequences of FIE are rather small, and mostly beneficial, largely because of the positive effects of fishing on growth. This prediction is made possible because of the crucial ecoevolutionary feedbacks among biomass, growth, and maturation, and the inclusion of growth as an evolving trait. Models that do not include these crucial factors might incorrectly predict a larger economic cost of evolution. Regardless, low fishing mortality is the key for successful management. Today, many fish stocks are still far from being managed in an ecologically optimal way. In such a case, our model predicts that FIE enables the stock to withstand higher harvests, but only if fishing mortality is not intermediately size-selective; otherwise, FIE may reduce economic yield and make the stock less viable. Admittedly, these evolutionary costs are small, but they may just be enough to push a fish stock from the state of overexploitation into collapse.

## Materials and Methods

Our bioeconomic model consists of two model components: the biological model, describing the life cycle of NEA cod, and the economic model, describing details such as cost and demand for the NEA cod trawl fishery. Each of

these components have been specifically estimated and calibrated for this stock by using data from 1932 to 2007 (Table S1). A more extensive model description can be found in *SI Materials and Methods: Model and Data Description*.

**Biological Model.** The biological model component is individual-based and has been developed in ref. 17, building on the ecogenetic modeling framework derived in ref. 1. The model describes each individual's growth, maturation, reproduction, and mortality in each year and follows the fate of ~50,000 superindividuals (37, 38). If a fish reproduces, genetic traits are inherited by offspring and expressed phenotypically. Mortality acts on these phenotypic traits, resulting in selection that may cause a genetic response in the life-history traits (Fig. 1A). We studied two versions of our model, an evolutionary and a nonevolutionary version, each modeling their respective population of individuals to compare a population that has the propensity to evolve with a population that does not evolve. We consider the evolution of four quantitative life-history traits: maturation tendency given by the (i) slope and (ii) intercept of a probabilistic maturation reaction norm (20), (iii) growth capacity, and (iv) reproductive investment. The genetic traits evolve independently, and we therefore do not account for pleiotropy or genetic linkage between traits. Our model has limitations, but thanks to the data availability for NEA cod, we are able to include estimates of the initial mean life-history trait values and annual exploitation rates, as well as parameters specifying the stock-recruitment relationship (describing fecundity and newborn mortality) and the density dependence of growth on stock biomass (17). Furthermore, a growth-survival tradeoff is included; the strength of this tradeoff was determined by matching the ecological properties for data on age and length at maturation, phenotypic growth, and biomass from 1932 to 1950 in the nonevolutionary version of the model after reaching demographic equilibrium (17). In the evolving population, the coefficient of genetic variation ( $CV_{z,G}$ ) has been determined empirically for each trait (17) by matching trends in age and length at maturation over a 74-y period (from 1932 to 2005). In this calibration, the historic selection pressure was mimicked by using annual harvest probabilities in the feeding and spawning grounds from 1932 until 2005. The resultant  $CV_{z,G}$  has been found to be lower than what was assumed in previous studies using the same modeling framework but not based on specific stocks (1, 16, 25, 26), as was the case here. For the non-evolving population, which is only driven by ecological processes, the  $CV_{z,G}$  is equal to zero.

**Economic Model and Harvest Control Rule.** The economic model specifies (i) the harvest function, (ii) the profit function, (iii) the procedure for allocating fishing quotas, and (iv) the demand function. All of these functions have been estimated and derived in detail in ref. 39 and used in ref. 22. We assume

a knife-edge selectivity (34, 35) that targets all fish above the size of 45 cm (17, 40). The biological and economic model components are linked through an annual feedback loop: SSB is fed into the economic model component where ultimately the TAC is determined by a HCR, and the derived TAC feeds back into the biological model component where it affects the stock size (Fig. 1, realized catch). The shape of the HCR is based on the one that has been implemented for NEA cod since 2004 (23, 24): the maximum fishing mortality  $F_{max}$  is allowed above a certain SSB level, given by the parameter  $B_{max}$ . Below  $B_{max}$ , fishing mortality decreases linearly to the origin (Fig. 1B). We explore model simulations over a large grid of combinations of  $F_{max}$  and  $B_{max}$ , searching for those combinations that achieve the economic objective of maximizing the net present value of fleet profits. All results, such as those for SSB and TAC, are given for a population that has been scaled up by a factor of 100,000. Because the model is stochastic, we ran each scenario for 15 independent replicates, and then averaged across these, presenting the mean in the tables and figures.

**Historic Fishing Pressure.** The observed harvest pressure in the feeding ground increased steadily from the 1930s to the middle of the 1960s and remained high until mid-2000. In the historic fishing scenarios, we use observed fishing mortalities from 1932 to 2005 and then assume a constant fishing mortality in the feeding ground ( $0.68 y^{-1}$ ) being maintained from 2006 into the future. This constant ( $0.68 y^{-1}$ ) is an average of the historic fishing mortality between 1946 and 2005 and is higher than what is considered to be precautionary for NEA cod ( $0.4 y^{-1}$ ) (40).

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# Supporting Information

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## SI Text

*SI Text* consists of two main sections: *SI Materials and Methods* and *SI Results*. In the first section, we describe the biological and economic components of the bioeconomic model, including a description of the data used to parameterize the model (Table S1). At the end of *SI Materials and Methods*, we discuss model limitations. In *SI Results*, we show in greater depth the emerging properties of the historic fishing scenario that may give rise to an evolutionary cost. Also, we investigate the implications of alternative discount rates for deriving optimal harvest control rules (HCRs). Furthermore, we probe into the robustness of our results. Simulating different levels of constant fishing mortality rates ( $0.2\text{--}0.8\text{ y}^{-1}$ ), we evaluate the impact of changing the minimum size limit, assuming a constant price, weight-dependent price, and, finally, changing the coefficient of genetic variation (i.e., evolvability) of the genetic life-history traits.

## SI Materials and Methods: Model and Data Description

**Biological Model.** The biological model is an individual-based model that uses the framework developed in ref. (1). This model combines quantitative genetics with ecological processes taking place at the individual level to derive knowledge on how fishing pressure progressively affects the stock at the population level. The genetic component of this model allows individuals to adapt to the selection pressure brought about by harvesting. The individual-based model follows about 50,000 superindividuals (2, 3). All models results, such as spawning stock biomass (SSB) and catch, are given for a population that has been scaled up by a factor of 100,000 to recreate realistic stock levels. Parameter values for our model (Table S1) are based on published sources, data collected by the Norwegian Institute of Marine Research (IMR), Knipovich Polar Research Institute of Marine Fisheries and Oceanography (PINRO), and the Norwegian Directorate of Fisheries, and survey data made available through the International Council for the Exploration of the Sea (ICES). This model has been developed and calibrated for the Northeast Arctic (NEA) cod stock in ref. 4. A similar model was used in ref. 5 for the same stock, without considering any evolutionary dynamics.

**Evolutionary dynamics.** This section describes first how we model the phenotypic expression of the genetic traits for individual maturation tendency, growth, and reproductive investment; second, how we introduce the distribution of the evolving genetic traits in the initial population; and finally, how the traits are inherited by offspring. Each genetic trait value  $z_G$  (denoted by subscript G) has a corresponding phenotypic trait value  $z_P$  (denoted by subscript P), with a genetic variance  $\sigma_{z,G}^2$  and phenotypic variance  $\sigma_{z,P}^2$ . At the population level, we assume phenotypic variance to be the sum of the genetic and environmental variance ( $\sigma_{z,P}^2 = \sigma_{z,G}^2 + \sigma_{z,E}^2$ ). Based on quantitative genetics (6), each trait has a heritability,  $h_z^2 = \sigma_{z,G}^2 / \sigma_{z,P}^2$ , which allows us to calculate the environmental variance  $\sigma_{z,E}^2 = \sigma_{z,G}^2 (h_z^2 - 1)$  for each trait in the initial population (where  $\sigma_{z,G}^2$  is empirically determined for each trait; see below). This environmental variance was then subsequently kept constant through time. The four considered quantitative genetic traits are the maturation tendency by a probabilistic maturation reaction norm (PMRN): (i) slope  $z_G = s_G$  and (ii) intercept  $z_G = i_G$ ; and (iii) growth capacity  $z_G = g_G$  and (iv) reproductive investment, given by the gonadosomatic index  $z_G = \text{GSI}_G$ . In the initial population, the genetic traits are assumed to be normally distributed with mean initial trait values and genetic variances determined by the coefficient of genetic variation  $\text{CV}_{z,G}$ , both based on empirical data (Table S1). The genetic traits are expressed

phenotypically by random draws from a normal distribution with means equal to the respective genetic trait (see Table S1 for initial values), with the corresponding environmental variances  $\sigma_E^2$ . We examined an evolutionary and a nonevolutionary version of the model, each modeling their respective population of individuals to compare a population that has the propensity to evolve with a population that does not evolve. First, the nonevolutionary model was calibrated to accomplish a match with data on NEA cod phenotypic growth, biomass, and age and length at maturation for the period 1932–1950 (4). For the nonevolving population, which is only driven by ecological processes, the  $\text{CV}_{z,G}$  by definition equals zero. In the evolving population,  $\text{CV}_{z,G}$  was determined by matching trends in age and length at maturation over a 74-y period (1932–2005); for all four evolving traits, these were then varied to determine the amount of evolution needed to match the maturation trends for 1932–2005. Based on previous models, the range of evaluated  $\text{CV}_{z,G}$  was between 0% and 12% (1, 7–9). All possible combinations were systematically evaluated and ranked by log-likelihood. The combination that ranked best was consequently selected and used to define the  $\text{CV}_{z,G}$  values for each trait.

Offspring inherited genetic trait values from their parents by drawing randomly from normal distributions with means equal to the midparental genetic trait values (i.e., the arithmetic mean trait value of the two parents) and variances equal to half the variance for a given genetic trait in the initial population (thus assuming a constant recombination–segregation–mutation kernel) (1, 10). After the initial year (e.g., the first year in the simulation), genetic means, heritabilities, and the trait distributions could change freely as determined by the processes of maturation, somatic growth, reproduction, natural mortality, and harvesting mortality. These processes were applied sequentially in each year to all individuals.

**Maturation, growth, reproduction, and mortality.** Each year, the probability  $p_m$  that an immature individual will mature is described by a PMRN (11, 12); this is a function of the individual's length  $l$  and age  $a$  and given by  $p_m = [1 + \exp(-(l - l_{p50,a})/\nu)]^{-1}$ . The length  $l_{p50,a}$  is where the maturation probability  $p_m$  is equal to 50% at age  $a$ , as given by  $l_{p50,a} = i_p + s_p a$ , with a phenotypic intercept  $i_p$  and slope  $s_p$ . The parameter  $\nu$  is determined by the lower bound probability  $p_l$  (25%) and the upper bound probability  $p_u$  (75%) of the maturation envelope (1, 4), together with the PMRN width  $w$ , as given by  $\nu = w / \ln \frac{p_l^{-1} - 1}{p_u^{-1} - 1}$ .

To reflect density dependence in growth brought about by changes in abundance, and consequently competition and resource availability, we used an estimated relationship of phenotypic growth  $g_{P,D,t} = g_{P,t} \exp(-x B_t)$ , depending on total stock biomass  $B_t$  in year  $t$ . The hypothetical length increment where biomass  $B_t$  is zero is referred to as the maximum growth increment, and  $x$  is the strength of density dependence reducing growth relative to this maximum. For this estimation (Table S1), derived in detail in ref. 4 and used in ref. 5, we used data on annual growth increments and biomass for the period 1978–2009, obtained from survey data and stock assessment (4, 13). The parameters were estimated by regressing log-transformed mean annual growth increments for ages 0–5 y in the winter survey against total biomass and other covariates ( $R^2 = 73\%$ ) (4). For the immature individuals, denoted by a superscript I, the body length in a given year depends on the length in the previous year and the growth increment in that year,  $l_t^I = l_{t-1}^I + g_{P,D,t-1}$ . Mature individuals, denoted by a superscript M, also allocate re-

sources to reproduction, depending on the reproductive investment; this is given by the phenotypic gonadosomatic index  $GSI_P$  and a conversion factor  $\gamma$ , needed to account for the higher energy content of gonadic tissue relative to somatic tissue (14, 15). Consequently, the length of a mature individual is given by  $l_t^M = 3(l_{t-1}^M + g_{P,D,t-1}) / (3 + \gamma GSI_{P,t-1})$ . An individual female's fecundity  $f$  is determined by its length  $l$  and gonadosomatic index phenotype  $GSI_P$  and given by  $f = kl^j GSI_P D$ , where  $D$  is the weight-specific packing density of oocytes (16), and  $k$  and  $j$  are allometric constants relating body length to body mass. The gonad weight at a given age can be calculated from fecundity by dividing it by the weight-specific packing density (shown in Fig. S1). An individual's probability to mate is proportional to its gonad mass, where large gonads due to larger body size and/or gonadosomatic index result in a higher production of gametes (eggs and sperm), and therefore in the production of more offspring. In our model, sex was assigned randomly at birth at a 1:1 primary sex ratio. Atlantic cod are batch spawners and so may mate with several different partners (17, 18). We therefore assumed mating to be random with replacement.

The individuals can die from natural or fishing mortality. In our model, natural mortality originated from three sources: newborn mortality, cost of growth, and a constant background natural mortality. The density-dependent newborn mortality was modeled by using an estimated Beverton–Holt stock-recruitment relationship (19) from virtual population analysis (VPA) data (20, 21). Recruitment depends on  $SSB_t$  in year  $t$  and sea surface temperature  $SST_t$ , reflecting the impact of climate. The sea surface temperature stretches from the Kola meridian transect (33°50' E, 70°50' N to 72°50' N) and has been shown to be a good indicator for recruitment for NEA cod (22–25). The expected number  $R_{3,t}$  of recruits at age 3 y is then given by  $R_{3,t+3} = c_0 SST_t + (c_1 SSB_t / (1 + c_2 SSB_t))$ , where  $c_0$ ,  $c_1$ , and  $c_2$  are statistically estimated parameters ( $R^2 = 58.9\%$ ). The two density-dependent parameters  $c_1$  and  $c_2$  were scaled to the modeled population (Table S1). Annual temperature data from 1932 to 2005 was fed into the modeled stock-recruitment relationship, and after 2006 we used the average from 1995 to 2005. In this stock-recruitment model, we ignore cannibalism, even though it has been shown to be important for natural mortality in young age classes (24, 26). We found the expected number  $R_{0,t}$  of newborn recruits by back-calculating the predicted number of 3-y olds, assuming an annual total natural mortality probability equal to  $0.2 y^{-1}$ , as conventionally done for this stock in assessment (21). The survival probability of the offspring of a given spawning pair was equal to  $R_{0,t}$  divided by the total fecundity of the spawning population.

The second source of mortality, the growth–survival tradeoff, accounts for less energy available for maintenance (27, 28) and lower survival as growth increases, which may be a result of, for example, risky foraging behavior (29, 30). We therefore included a tradeoff between an individual's survival and genetic growth capacity  $g_G$  through the extra mortality probability  $m_g = g_G / g_{max}$ , where  $g_{max}$  is the maximal genetic growth increment at which the survival probability drops to zero, and determines the strength of this tradeoff. The parameter  $g_{max}$  is a priori unknown and has been determined in a nonevolutionary model to imitate the stock demographically from 1932 to 1950 (4), by varying  $g_{max}$  from 50 to 200 cm, in steps of 5 cm, resulting in 31 evaluated combinations. This grid covered the range of values being assumed in published versions of this model (1, 7). Comparing model predictions with time-series data on phenotypic growth, biomass, and mean age and length at maturation for the period 1932–1950, the growth–survival tradeoff  $g_{max}$  was determined by log-likelihood (Table S1). Together, the background natural mortality and the additional mortality resulting from the growth–survival tradeoff produced annual natural mortality probabilities  $m$  equal to 0.18, as assumed by ICES in its VPA analyses (Table S1).

As is the case for NEA cod, harvesting was implemented in the model separately in the feeding grounds and spawning grounds. In the feeding grounds, harvesting was size-selective with minimum size limits within the range recorded for NEA cod from the 1980s onward (31). In the spawning grounds, only mature individuals were harvested, and there was no minimum size limit. Due to annual spawning migration out of the feeding grounds for approximately one-quarter of the year, the harvest probability of mature fish on the feeding grounds was  $1 - (1 - p_0)^{3/4}$ , where  $p_0$  is the harvest probability for the immature fish.

**Economic Model.** To calculate the welfare effects of harvesting, we first specify the harvest function; second, specify the profit function; third, derive a procedure for allocating fishing quotas; and fourth, derive the demand function. All of these functions have been estimated and derived in detail in ref. 32 and used in ref. 5. Furthermore, we specify the objective functions to derive an optimal HCR.

**Harvest function.** Following refs. 33 and 34, the harvest function of vessel  $i$  in year  $t$  is given by a Cobb–Douglas production function  $h_{it} = q B_t^\alpha e_{i,t}^\beta$ , where  $q$  is a catchability coefficient,  $B_t$  is the amount of total stock biomass, and  $e_{i,t}$  is fishing effort. In our model, effort is defined as the number of days a boat is fishing cod north of 62° N, multiplied by the size (given in gross tonnage) of the boat. The stock–output elasticity  $\alpha$  and effort–output elasticity  $\beta$  describe how harvest changes when the respective inputs, biomass, and effort change.

**Profit function.** The cost data for each vessel contains expenses made for labor wages and shares to crew; social expenses (i.e., payroll-related expenses, such as employer contributions to pension and the employer portion of social security tax); fuel and lubrication oil; bait, ice, salt, and packaging; food expenses to crew, as well as maintenance on vessel, maintenance and investment on gear, insurance on vessel, other insurances, depreciation on vessel, and other operating expenses (35). In total, there are 11 cost components, which are indexed  $k = 1 \dots 11$ . Total costs incurred by vessel  $i$  in year  $t$  are given by the vector of nominal cost components  $C_{i,k,t}$ , which are subsequently corrected for inflation using the Producer Price Index (PPI). We calculate the part of the total costs incurred for catching cod by the share of days vessel  $i$  spends on catching cod in the total number of days vessel  $i$  is fishing at sea. Using index  $j$  to enumerate all eight fish species caught (with cod being  $j = 8$ ) and denoting the number of days in year  $t$  that vessel  $i$  catches species  $j$  by  $D_{ij,t}$ , the total number of days vessel  $i$  spends catching fish in year  $t$  is equal to  $\sum_{j=1}^8 D_{ij,t}$ . Therefore, the costs attributed to catching cod by vessel  $i$  in year  $t$  are  $C_{i,t} = (D_{i8,t} \sum_{k=1}^{11} C_{i,k,t}) / (PPI_t \sum_{j=1}^8 D_{ij,t})$ .

We empirically determine which fraction of the costs of fishing per boat  $C_{i,t}$  comprise fixed and variable costs by estimating  $C_{i,t} = c_f + c_v e_{i,t}$ , where  $c_f$  can be interpreted as fixed costs, and  $c_v$  are variable costs. Multiplying the catch  $h_{i,t}$  of vessel  $i$  with the price of cod  $P_t$  yields the revenue  $P_t h_{i,t}$  of vessel  $i$ . The profit  $\pi_{i,t}$  of vessel  $i$  is then given by offsetting this revenue with the costs of vessel  $i$  and given by  $\pi_{i,t} = P_t h_{i,t} - c_f - c_v e_{i,t}$ .

**Issuing individual quotas.** Harvest quotas could in principle be allocated through a market mechanism, such as an auction, or handed out by the government to the boat owners. It is not clear a priori what the most efficient allocation (or market outcome) is, because the size of the quota and number of quotas can vary. Each boat faces a fixed cost, but is harvesting less efficiently when the size of the quota per boat increases, determined by the estimated effort–output elasticity (parameter  $\beta$  in Table S1). For each year  $t$ , we identify an optimal number  $n_t^*$  of vessels harvesting an optimal number  $e^*$  of tonnage days for a given total allowable catch (TAC) and total stock biomass (for details, see ref. 32), where  $n_t^* = H_t q^{-1} e^{*-\beta} B_t^{-\alpha}$ .

**Demand function.** The NEA cod fishery contributes a large part of the world's cod landings and therefore affects the international

market price for cod. To describe this relationship, we use a linear demand function,  $P_t = b_0 - b_1 H_t$ , where  $P_t$  is the price for cod in year  $t$ ,  $H_t$  is the total harvested biomass in year  $t$  (as determined by the TAC), and  $b_0$  and  $b_1$  are parameters. The inverse price elasticity is estimated to be 0.5, i.e., if the supply of cod increases by 1%, the world price drops by 0.5% (32). Using the average kilogram price in the period 1998–2007 [in 2000 Norwegian kroner (NOK)] of 12.59 NOK, and the average landing of 527,800 tonnes allows us to solve for  $b_0$  and  $b_1$  (Table S1).

**Objective function and HCR.** Each year, the NEA cod fishery generates economic profits for the fishing fleet, given by  $\Pi_t$ . Finding the maximum economic yield requires us to maximize the NPV of the fishery over  $T$  years, as given by  $\text{NPV} = \sum_{t=0}^T \Pi_t / (1+\delta)^t$ , where  $\delta$  is the discount rate.

The HCR implemented for the NEA cod fishery in 2004 translates precautionary reference points into a management plan (21, 36). Below these reference points, the stock is at risk for being harvested unsustainably. The implemented HCR for the NEA cod in 2004 consists of two parameters (37, 38): a maximum fishing mortality  $F_{pa}$  is followed if the spawning stock biomass level is above the precautionary biomass level  $B_{pa}$ ; below this biomass level, the fishing mortality decreases linearly to the origin, i.e., fishing mortality is zero at a biomass level of zero.

Here, we generalize a HCR with two parameters (Fig. 1B), which can be compared with the implemented management plan. If the SSB is between zero and  $B_{max}$ , the instantaneous fishing mortality for the given year is given by  $F_{max} \text{SSB} / B_{max}$ . If the SSB is larger than  $B_{max}$ , the fishing mortality is equal to  $F_{max}$ . The current HCR is therefore recovered as a special case when  $B_{max} = B_{pa}$  and  $F_{max} = F_{pa}$ . In our model, we vary the parameters in the HCR over a wide range of values, not constraining them to existing precautionary reference points. We search for the combination of parameter values  $B_{max}$  and  $F_{max}$  that deliver the best results for the objective function (maximize the net present value of fleet profits) and identify those as optima. The grid for the parameters covered 4,141 different HCRs. The parameter  $B_{max}$  was varied from 0 to 800,000 tonnes in steps of 20 tonnes, and the instantaneous fishing mortality  $F_{max}$  was varied from 0.2 to  $1.2 \text{ y}^{-1}$  in steps of  $0.01 \text{ y}^{-1}$ . Our model is individual-based, and for some of these HCRs, fishing could make the abundance very low. To avoid stochastic effects at low abundances, we therefore set a threshold below which the population was classified as extinct (at 20 modeled mature superindividuals) (3, 4). The computations were completed on Abel, a computer cluster with 10,000+ cores at the Research Computing Services at the University of Oslo.

**Model Limitations.** As with all models, our bioeconomic model has limitations and involves simplifications. A few assumptions merit special attention here. First, we assume an initial 1:1 sex ratio, although it has been shown that the sex ratio has fluctuated over time in this cod stock (39). Second, we assume no sexual selection, although it is possible that sexual selection may influence the evolutionary changes in life-history traits (40–42). Third, we do not include genetic correlations between the life-history traits describing maturation tendency, growth capacity, and reproductive investment (4). Fourth, we assume a constant minimum size limit that determines the harvestable biomass (Table S1), implicitly assuming knife-edge selectivity (19, 43), which may not be fully realistic. Although our size limit is based on data, the size selectivity has varied over the considered time period and across vessels since 1932 (for a sensitivity analysis with respect to minimum size limit, see Table S5). Fifth, the shape of the HCR we are considering is constrained by two parameters, reflecting the current management plan. Investigating completely different shapes or considering HCR parameters that change over time is an interesting avenue for further research. Sixth, we focused on the fishery

in the stock's feeding grounds and kept the fishing mortality at observed levels in the stock's spawning grounds; we did this because we wanted to mimic the historic selection pressure on the mature fish, while parsimoniously asking what can be changed for the trawler fleet in the Barents Sea. This assumption could be changed, and the next step would be to derive an optimal HCR for each of these fisheries.

## SI Results

**Historic Fishing Pressure.** Table S4 shows the harvesting properties for the scenario of historic fishing pressure (i.e., high fishing mortality) presented in Fig. 2. The evolutionary model delivers lower TAC, total biomass from age 3 y, and lower NPV, whereas the SSB is slightly higher compared with the nonevolutionary model.

In Fig. S1 we show the life-history changes in the scenario of historic fishing pressure, corresponding to Figs. 2 and 3. Genetic adaptations caused by fishing pressure lead to higher reproductive investment (Fig. S1A) and genetic growth (Fig. S1B). As a result, the evolving population has consistently larger gonad weight (Fig. S1C) and higher phenotypic growth (Fig. S1D). Due to evolutionary changes, the ratio between spawning stock biomass and total biomass changes over time because of a change in maturation schedule (Fig. S4), and this may have implications for stock assessment and the target reference points that are used for management.

**Alternative Discount Rates.** Table S2 presents optimal HCR derived for alternative discount rates. As expected, higher discount rates lead to slightly higher fishing mortality, even though only marginally. This finding may seem surprising, but happens because larger catches result in lower prices, and hence profits. At a certain point, the resulting profit loss from lower prices outweighs the profit gain resulting from catching more fish, irrespective of the discount rate (5).

**Alternative Scenarios with Constant Fishing Mortalities. Constant and weight-dependent prices.** We probe the robustness of our results by varying the fishing mortality under alternative assumptions and investigating how this influences the effects of evolutionary changes. First, as a theoretical exercise, we assume that sales prices are independent of the total catch and the price is constant; this is clearly not realistic for the NEA cod fishery, but certainly the case for many other fisheries. As a constant price, we use the inflation-corrected average kilogram price in the period 1998–2007 of 12.59 NOK. Second, in addition, we assume that sales prices are weight-dependent, i.e., the price that can be obtained per kilogram of cod rises with the weight of the fish; we found little evidence that this is actually the case for the fleet of trawlers we are considering here, but it may be relevant for other vessel types, notably smaller coastal vessels. As a theoretical benchmark, we can rely on the minimum prices from the Norwegian fishermen's sales organization (44). The prices for the different weight classes are as follows. Cod that is heavier than 6.5 kg yields 17 NOK per kilogram. Cod that weighs between 2.5 and 6.5 kg yields 14.25 NOK per kilogram; cod that weighs between 1.0 and 2.5 kg yields 12.25 NOK per kilogram; and all cod that weighs less than 1.0 kg yields 9.25 NOK per kilogram. Table S3 shows the emerging properties of different fishing mortalities and the NPV for a constant price ( $\text{NPV}_{CP}$ ) and for weight-dependent prices ( $\text{NPV}_{WP}$ ). For comparison, we also show the NPV derived from the model used in the main text. We find that our earlier results presented in Table 1 fully carry over to the case where the price is constant or weight-dependent. Still, evolution increases the NPV of a fishery if fishing mortality is low, and it decreases the NPV of a fishery if fishing mortality is high (Table S4).

**Emerging properties for minimum size limits of 25, 45, and 85 cm.** Table S5 shows the emerging properties of TAC, total biomass above the age of 3 y, and NPV that complement Fig. 4. For a fishing mortality of  $F = 0.8 \text{ y}^{-1}$  and a minimum size limit equal to 25 cm, both the evolving and the nonevolving model population go extinct. When harvest pressure is high and the size limit is low (25 cm), the economic losses due to evolution that we see at a size limit of 45 cm disappear. The NPV values are overall lower, however, for the 25-cm size limit than for the 45-cm size limit. At a very high minimum size of 85 cm, the nonevolutionary model performs insignificantly better than the evolutionary model, suggesting that economic losses from evolutionary change are not increasing as minimum sizes increase. Instead, those evolutionary costs are highest (albeit still small) for a minimum size of 45 cm—a size that is based on historic values for the Norwegian and Russian cod fisheries (4, 21) and very close to the size currently used as a legal minimum size (45). Fig. S2 shows the final genetic trait values (year 2100) for different fishing mortalities ( $F = 0.4, 0.8 \text{ y}^{-1}$ ) and for different minimum size limits. We find here that the evolutionary change is larger as fishing mortality increases for all traits except for growth, and that higher minimum size limits result in lower selective pressure and less evolution (Fig. S2).

**Varying the coefficients of genetic variation and fishing mortality.** In the evolutionary version of our model, the coefficient of genetic variation (Table S1) has been determined empirically by matching trends in age and length at maturation over a 74-y period (1932–2005) (4). The genetic changes emerging from this study are

found to be lower than what has been predicted in comparable studies (1, 7–9). These studies assumed a coefficient of genetic variation for all traits equal to 8% and 6%, respectively. As a robustness check, we therefore used these higher coefficients of genetic variation (6% and 8%) and performed simulations for different fishing mortalities. After fishing with a particular fishing mortality from 1932 to 2100, we compare the simulation end-points for age at maturation and TAC with our calibrated evolutionary model (Table S1) and nonevolutionary model (for which all coefficients of genetic variation are equal to zero). As expected, we find that an assumed  $\text{CV}_{z,G}$  of 6% and 8% results in predicting larger evolutionary responses, causing a much lower age at maturation in the year 2100 (Fig. S3A). As genetic variance increases, the fish mature at a younger age and at a smaller size, and also grow faster. As fishing mortality increases, age at maturation also declines for the nonevolutionary model, which is entirely due to phenotypic plasticity and density dependence in response to a lower abundance of the stock. For the case where the coefficients of genetic variation are set to 6% and 8%, we find that for a given fishing mortality, higher TACs can be obtained compared with the models where evolutionary change is weaker or even absent. This finding indicates that stronger evolutionary forces tend to have as positive effect on the TAC, and corroborates our earlier finding that evolution toward faster growth tends to have positive effects on the fishery.

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**Table S1. Parameter values and data sources for the bioeconomic model**

Parameters	Value	Source
<b>Biological model component</b>		
Initial mean PMRN slope, $\bar{s}_G$	0.15 cm y <sup>-1</sup>	4, 12
Initial mean PMRN intercept, $\bar{i}_G$	77.4 cm	4, 12
Initial mean reproductive investment, $\overline{GSI}_G$	0.15	17
Initial mean growth capacity, $\bar{g}_G$	11.08 cm	4, 5, M. Heino*
PMRN width, $w$	12.88 cm	4, 12
Coefficient of genetic variation in PMRN slope, $CV_{s,G}$	10%	4
Coefficient of genetic variation in PMRN intercept, $CV_{i,G}$	2%	4
Coefficient of genetic variation in reproductive investment, $CV_{GSI,G}$	12%	4
Coefficient of genetic variation in genetic growth, $CV_{g,G}$	4%	4
Initial heritability, $h_z^2$	0.2	6
Strength of density dependence in growth, $x$	2.08 10 <sup>-5</sup> kg <sup>-1</sup>	4, 5, M. Heino*
Reproductive investment conversion factor, $\gamma$	0.60241	14
Allometric constant, $k$	3.2 10 <sup>-6</sup> kg cm <sup>-j</sup>	O. S. Kjesbu*
Allometric exponent, $j$	3.24	O. S. Kjesbu*
Weight-specific oocyte density, $D$	4.45 10 <sup>9</sup> kg <sup>-1</sup>	16
Maximal growth capacity, $g_{max}$	105 cm	4
Stock recruitment constant, $c_1$	0.7549 kg <sup>-1</sup>	4, 20, 21
Density-dependent stock recruitment constant, $c_2$	-6.0633 kg <sup>-1</sup>	4, 20, 21
Temperature coefficient in stock recruitment, $c_0$	0.4241 °C <sup>-1</sup>	PINRO**, 22, 23
Natural mortality probability, $m$	0.18	21
Immature fishing probability in spawning grounds pre-1932	0.38	M. Heino*, O. R. Godø*
Immature fishing probability in feeding grounds pre-1932	0.09	M. Heino*, O. R. Godø*
Minimum size limit on feeding grounds	45 cm	21, O. R. Godø*
<b>Economic model component</b>		
Intercept of the demand function, $b_0$	18.88 NOK kg <sup>-1</sup>	32
Slope of the demand function, $b_1$	1.19 10 <sup>-8</sup> NOK kg <sup>-2</sup>	32
Stock-output elasticity, $\alpha$	0.58	32
Effort-output elasticity, $\beta$	0.85	32
Catchability coefficient, $q$	6.17 10 <sup>-4</sup> tonnes <sup>-1</sup> d <sup>-1</sup>	32
Fixed costs per boat, $c_f$	1.55 10 <sup>6</sup> NOK	32
Variable costs per boat, $c_v$	131.6 NOK tonnes <sup>-1</sup> d <sup>-1</sup>	32
Optimal number of tonnage days, $e^*$	66,712 tonnes/d	32

Economic data for the Northeast Arctic cod fishery: costs and harvests from the Norwegian Directorate of Fisheries, Bergen, Norway, provided by P. Sandberg; biomass and total landings are from ICES (21); and demand function is from Statistics Norway, Oslo, Norway, and Norwegian Directorate of Fisheries. Economic values have been inflation corrected using the Producer Price Index from the Organisation for Economic Cooperation and Development, Paris, France, with year 2000 as a baseline. The applied exchange rate is 1 US dollar = 5.6 NOK. Biological data for the Northeast Arctic cod stock is described below. \*IMR, Institute of Marine Research, Bergen, Norway: Survey data on growth from 1932 to 2009 provided by M. Heino; allometric data from survey 1999–2007 provided by O. S. Kjesbu; data on fishing mortality and minimum size limit provided by M. Heino and O. R. Godø. \*\*PINRO, Knipovich Polar Research Institute of Marine Fisheries and Oceanography, Murmansk, Russia: Temperature data.

**Table S2. Optimal HCR with parameters  $F_{max}$ ,  $B_{max}$ , and corresponding NPV for different discount rates ( $\delta$ ), 0%, 2%, and 4%**

Model	$\delta$ , %	$F_{max}$	$B_{max}$	$F$	TAC	SSB	NPV
Evolution	0	0.33	100	0.33	467 (60)	801 (163)	96.0
	2	0.34	20	0.34	469 (60)	767 (163)	25.4
	4	0.35	20	0.35	470 (60)	735 (155)	12.6
Ecology	0	0.33	40	0.33	439 (48)	670 (125)	94.7
	2	0.35	100	0.35	443 (48)	643 (118)	25.3
	4	0.36	100	0.36	445 (48)	618 (114)	12.6

Averages of fishing mortality ( $F$ ), total allowable catch (TAC), and spawning stock biomass (SSB) with temporal SDs in parentheses. Units:  $F_{max}$  and  $F$  (y<sup>-1</sup>);  $B_{max}$ , TAC, and SSB (1,000 tonnes); NPV (in billions, US dollars). HCR, harvest control rule; NPV, net present value.

